

Definition of Random Variable

In simple terms, a random variable (also referred to as a stochastic variable) is a real-valued set function whose value is a real number determined by the outcome of an experiment. The range of a random variable is the set of all the values it can assume. More formally, in measure theoretic terms, a random variable X is a real-valued function that maps \mathcal{S} into \mathcal{R} and satisfies the condition that for every Borel set $B \in \mathcal{B}$ the inverse image $X^{-1}(B) \in \mathcal{F}$, where $X^{-1}(B) = \{s : s \in \mathcal{S} \text{ and } X(s) \in B\}$

Definition of Cumulative Density Function

The real-valued function $F(x)$ such that $F(x) = P_x\{(-\infty, x]\} = P(X \leq x)$ for each $x \in \mathcal{R}$ is called the distribution function, also known as the cumulative distribution/density function, or CDF.

Definition of Probability Density Function

For a random variable X if there exists a nonnegative function $f(x)$, defined on the real line, such that for an interval B , $P(X \in B) = \int_B f(x)dx$, then X is said to have a continuous distribution and the function $f(x)$ is called the probability density function or simply the density function (or PDF).

Moreover, for a continuous random variable, we can find the following properties

$$(i) \quad F(x) = \int_{-\infty}^x f(u)du$$

$$(ii) \quad f(x) = F'(x)$$

$$(iii) \quad \int_{-\infty}^{\infty} f(u)du = 1$$

$$(iv) \quad F(b) - F(a) = \int_a^b f(u)du$$

Definition of Normal Distribution

Normal distribution has the following density:

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right], \quad x \in \mathcal{R}$$

The distribution is written as $X \sim N(\mu, \sigma^2)$. The special case of the normal distribution for $\mu=0$ and $\sigma=1$ is known as the standard normal and it has the following properties:

$$(i) \quad E(X) = 0$$

- (ii) $V(X) = 1$
- (iii) $E(X^3) = 0$
- (iv) $E(X^4) = 3$

Definition of Lognormal Distribution

The Lognormal distribution can be extended by letting $Y = \ln X \sim N(\mu, \sigma^2)$. X has the following PDF

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{(\ln x - \mu)^2}{2\sigma^2}\right] \frac{1}{x}, \quad x > 0$$

Furthermore, we have:

- (i) $E(X) = e^{\mu + \sigma^2/2}$
- (ii) $V(X) = e^{2\mu} [e^{2\sigma^2} - e^{\sigma^2}]$

Definition of Convergence in Distribution

Given a sequence of random variables X_n whose CDF is $F_n(x)$, and a CDF $F_X(x)$ corresponding to the random variable X , we say that X_n converges in distribution to X , and write $X_n \xrightarrow{d} X$, if $\lim_{n \rightarrow \infty} F_n(x) = F_X(x)$ at all points x at which $F_X(x)$ is continuous.

Definition of Convergence in Probability

The sequence of random variables X_n is said to converge in probability to the real number x if $\lim_{n \rightarrow \infty} P[|X_n - x| > \varepsilon] = 0$ for each $\varepsilon > 0$. Thus it becomes less and less likely that the random variable $X_n - x$ lies outside the interval $(-\varepsilon, +\varepsilon)$. The sequence of random variables X_n is said to converge in probability to the random variable X if the sequence of their difference $X_n - X$ converges in probability to 0.

We write $X_n \xrightarrow{p} X$.

Definition of Convergence in mean-squared

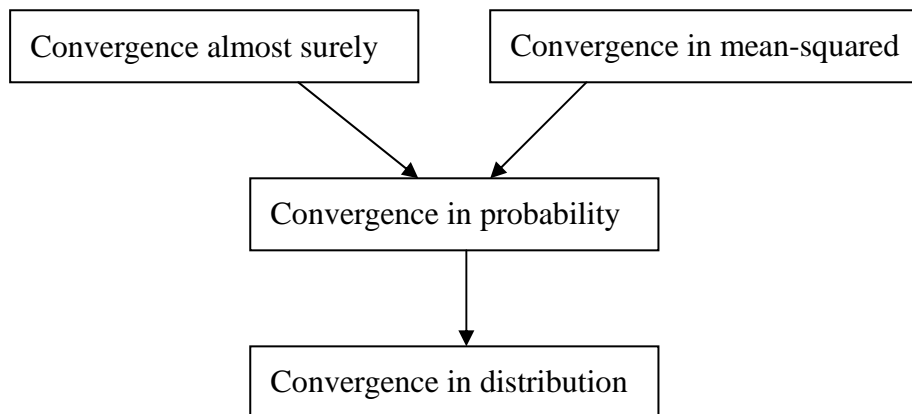
The sequence of random variables X_n is said to converge in mean-squared to X and designated $X_n \xrightarrow{(2)} X$ or $X_n \xrightarrow{M.S.} X$, if $E[|X_n - X|^2]$ exists and $\lim_{n \rightarrow \infty} E[|X_n - X|^2] = 0$, that is, if the 2nd moment of the difference tends to zero.

Definition of Convergence Almost Surely

The sequence of random variables X_n is said to convergence almost surely to the real numbers x , and is written as $X_n \xrightarrow{a.s.} x$, if $P[\lim X_n = x] = 1$. In other words, the sequence X_n may not converge everywhere to x , but the points where it does not converge form a set of measure zero in the probability sense. X_n is said to converge almost surely to the random variable X if $X_n - X \xrightarrow{a.s.} 0$.

Relationships among Modes of Convergence

We have the following relationships among the above convergence modes (the black arrows from A to B mean that “A implies B”):



Counter examples:

- (i) Suppose the sample space is $S = [0,1]$ and X_n is a sequence of random variables defined as follows:

$$X_n(s) = \begin{cases} 1, & 0 \leq s \leq \frac{1}{2} \\ 0, & \frac{1}{2} < s \leq 1 \end{cases}$$

With a probability of one-half for each case. The corresponding sequence of distribution functions is

$$F_{X_n}(s) = \begin{cases} 0, & s < 0 \\ \frac{1}{2}, & 0 \leq s \leq \frac{1}{2} \\ 1, & s > \frac{1}{2} \end{cases}$$

Next suppose the random variable X is defined as

$$X(s) = \begin{cases} 0, & 0 \leq s \leq \frac{1}{2} \\ 1, & \frac{1}{2} < s \leq 1 \end{cases}$$

with a probability of one-half for each case. It's easily verified that the

corresponding CDF is the same as that for X_n , which means that $X_n \xrightarrow{d} X$.

However, $X_n - X$ is always equal to 1 and hence X_n does not converge in probability to X .

(ii) $X \sim U[0,1]$ and binary intervals $I_1 = [0,1], I_2 = [0, \frac{1}{2}], \dots$ so that

$$I_{2^m+i} = [\frac{i}{2^m}, \frac{i+1}{2^m}]$$

for $m=0,1,2,\dots$ and $i=0,1,\dots, 2^m-1$. Therefore, the 2^m intervals of length 2^{-m} cover the interval $[0,1]$.

Let $Y_n = 1$ if $X \in I_n$ and $Y_n = 0$ if $X \notin I_n$. The sequence Y_1, Y_2, \dots converges in probability to $Y=0$ since, for all $0 < \varepsilon \leq 1$, $P(|Y_n - 0| \geq \varepsilon) = \Pr(Y_n = 1) = P(X \in I_n) = \text{length of interval } I_n \rightarrow 0, \text{ when } n \rightarrow \infty$. Y_n does not converge almost surely to zero since any X is in only one of the 2^n intervals of length 2^{-n} , for all n . In other words, for all $\omega \in \Omega$, $Y(\omega)$ assumes the value 1 for an infinite number of n 's and assumes the value 0 for an infinite number of n 's. Hence, for each $\omega \in \Omega$, $Y(\omega)$ does not converge.

(iii) Let $X_n = \begin{cases} 0, p = 1 - \frac{1}{n} \\ n, q = \frac{1}{n} \end{cases}$ hence we have $X_n \xrightarrow{p} X$. However,

$$E[(X_n - 0)^2] = n^2 \frac{1}{n} \rightarrow \infty, \text{ as } n \rightarrow \infty.$$

(iv) Let $\omega \sim U(0,1)$, and $X_n(\omega) = n1_{(0, \frac{1}{n})}(\omega)$. $\forall \omega \in [0,1]$ and $\varepsilon > 0$, there

$\exists N_0$ s.t. $n > N_0$ with $\frac{1}{n} < \omega$ and $|X_n(\omega) - 0| < \varepsilon$. It implies that

$X_n \xrightarrow{a.s.} 0$. However, $E[|X_n - 0|^2] = 1 \not\rightarrow 0$. Thus, X_n does not converge to 0 in mean-squared sense.

The Ito Integral

We are sufficient to prove $\sum (dw)^2 \xrightarrow{M.S.} t$.

Let $Q = \sum (dw)^2 = \sum \Delta z^2$, where $dw = \sqrt{\Delta} z$ $z \sim N(0,1)$ And $\Delta = \frac{t}{n}$

$$E[Q] = t$$

$$\text{Var}(Q) = n \text{Var}(\Delta z^2) = n \Delta^2 \text{Var}(z^2)$$

$$\begin{aligned} \text{Var}(z^2) &= E(z^4) - E^2(z^2) \\ &= 3 - 1 = 2 \end{aligned}$$

$$\therefore \text{Var}(Q) = 2n\Delta^2 = 2n \left(\frac{t}{n}\right)^2 \xrightarrow{n \rightarrow \infty} 0$$

Because $\text{Var}[Q] = E((Q-t)^2) \xrightarrow{n \rightarrow \infty} 0$, $\sum (dw)^2 \xrightarrow{M.S.} t$